

Trivariate Models for Stochastic Episodes with Applications to Hydrology, Climate and Finance

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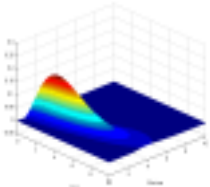
University of Nevada Reno, USA

with

Tomasz J. Kozubowski and Fares Qeadan

“From the tropics to the arctic, climate and weather have powerful direct and indirect impacts on human life. Extremes of heat and cold can cause potentially fatal illnesses, Other weather extremes, such as heavy rains, floods, and hurricanes, also have severe impacts on health. Approximately 600 000 deaths occurred worldwide as a result of weather-related natural disasters in the 1990s.”

The WHO, 12 November, 2008



Outline



- **Motivation:** Extreme hydroclimatic and weather events, e.g. drought, flood, or heat waves; Financial modeling.
- **Examples:** (1) Paris heat wave, August 2003.
 - (2) Great Drought of the 1930s in the USA, THE “DUST BOWL”,
 - Financial modeling – stability “with respect to growth period”
- **Stochastic Models:** First joint model for duration N , magnitude X and maximum Y of events: (N, X, Y)
- **Mathematical Results:** The model for (N, X, Y) :
 - Stochastic representation, pdf;
 - Marginals: one dimensional and bivariate; conditionals;
 - Correlation structure
 - Distribution of the ratio max/sum.
- **Examples:** Back to the “DUST BOWL”, heat wave, and financial modeling.
- **Summary**

Motivation: Climate and hydrology - drought

"DUST BOWL" - Great drought of the 1930s in the USA, the setting of John Steinbeck's "Grapes of Wrath".



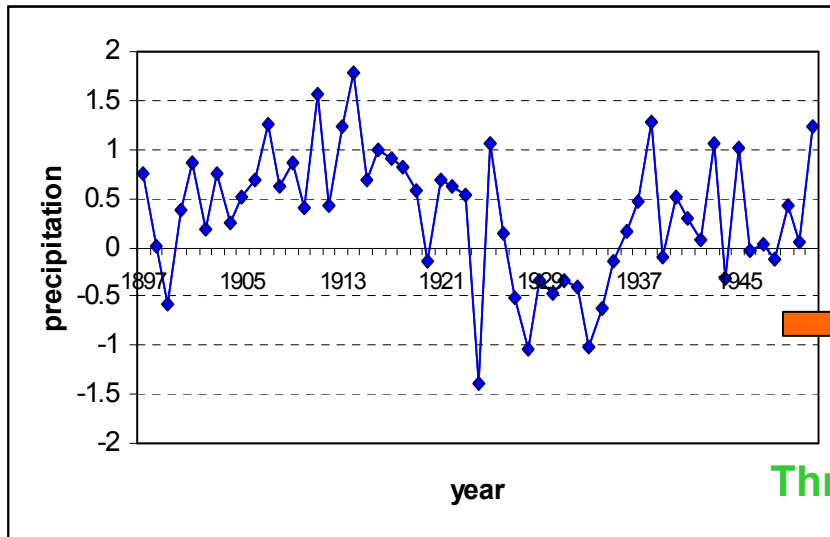
What are the chances of a drought with duration exceeding 10 years?

What are the chances that "Dust Bowl" will happen again?

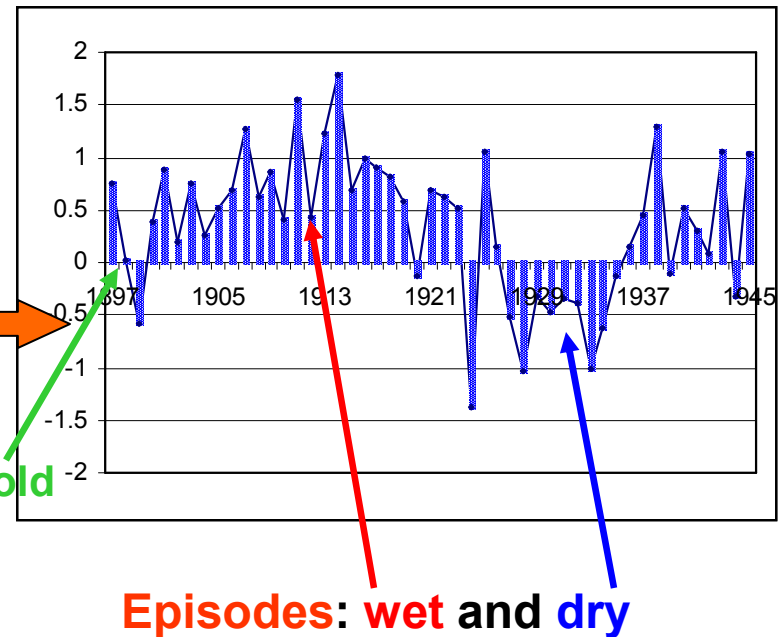
What are the chances that a 5 years long drought will have magnitude exceeding "Dust Bowl's" and a large maximum?

Motivation – drought, contd.

Original precipitation data in standard deviation units plotted as difference from a threshold.



Precipitation “events” or “episodes”.

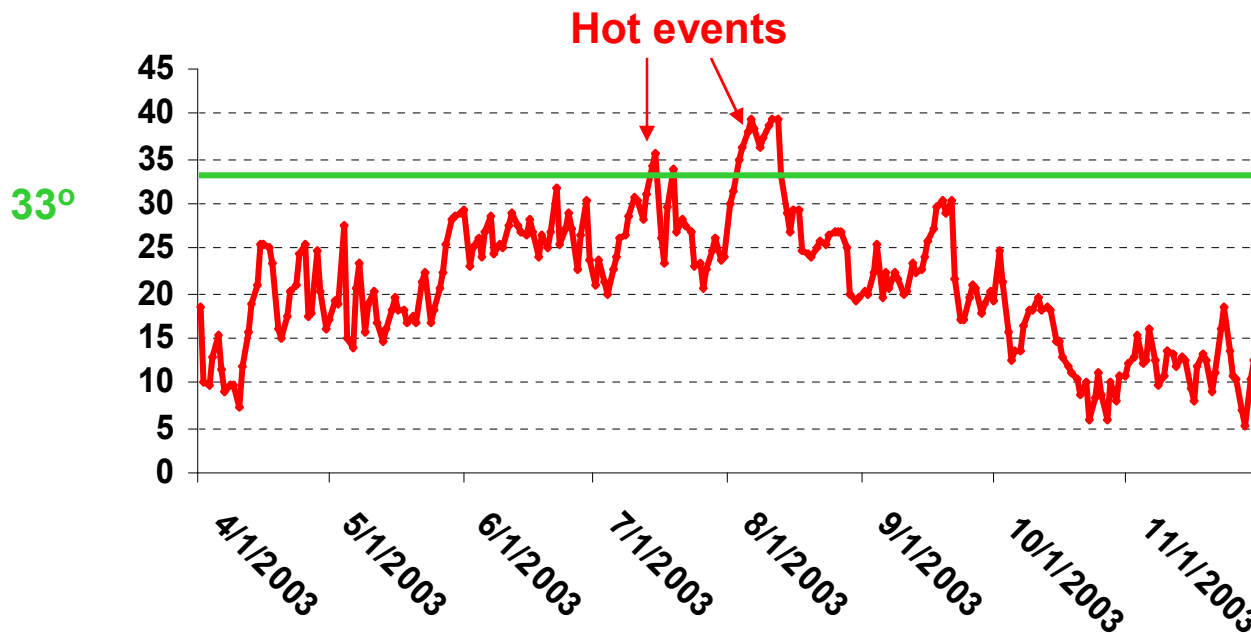


An **episode** is a period with the process staying consecutively above/below threshold: e.g. “dry”, “wet” year, drought, flood, etc.

Threshold for “dry” or “wet” depends on the definition of the episode (e.g. drought).

Motivation: Climate and weather – heat waves

- In August 2003, France experienced an extreme heat wave, that resulted in an estimated 14,802 deaths*.
- Hot event: consecutive observations above the 33°C.



What are the chances that a large heat wave will happen again?

What are the chances of a hot event with duration equal to this one, 11 days?

What are the chances that a heat wave longer than 6 days and larger than current 98th percentile of heat waves' magnitudes?

*Dhainaut et al. 2004.

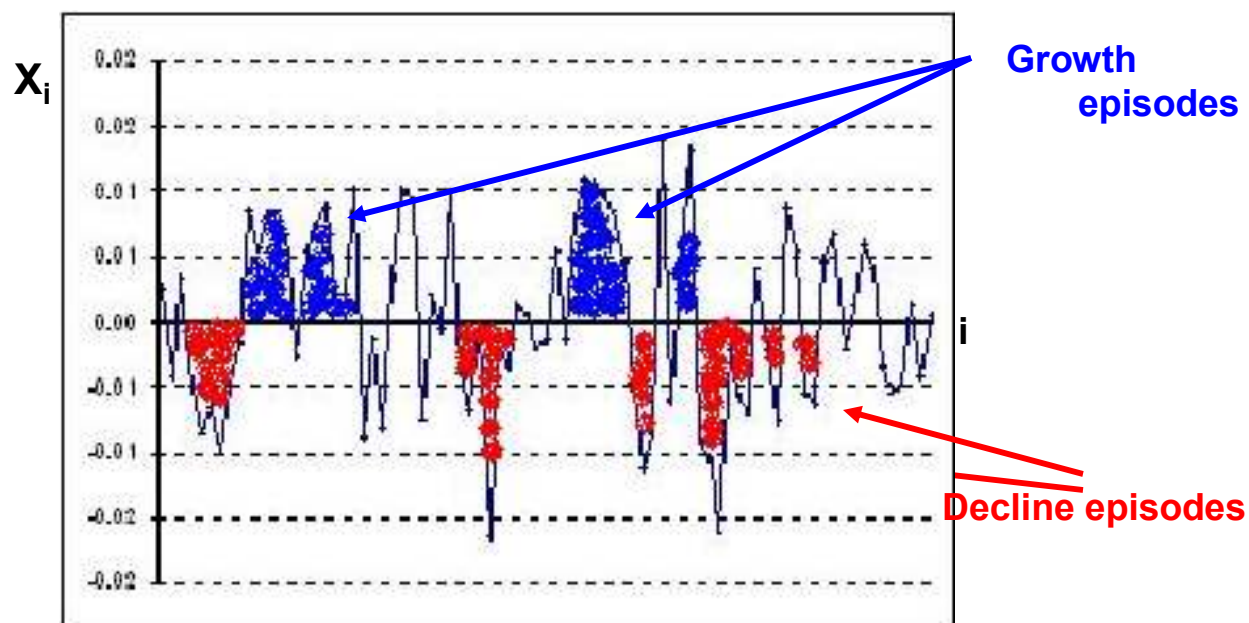
Motivation: Financial growth/decline episodes

Daily exchange rates between Japanese Yen and British pound quoted in UK pound, Jan. 2, 1980-May 21, 1996.



Process X_i : Daily log returns $n=4274$,
 $X_i = \log(\text{Rate_day_i} / \text{Rate_day_i-1})$

Episodes: consecutive days when the exchange rates were growing/declining, i.e.:
Growth $X_i > 0$, Decline $X_i < 0$.

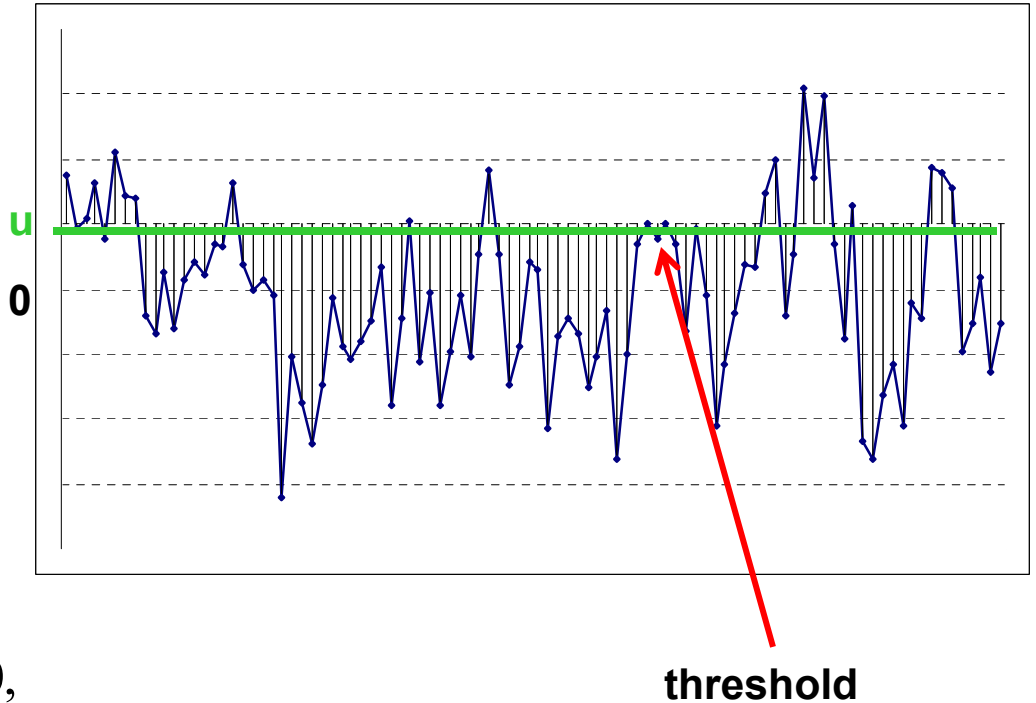


N = length of a growth period, $X = \sum X_i$ = cumulative log return over a growth period,
 Y = maximum log return over the growth period.

Motivation – theoretical: Peak over Threshold theory

Take a process X_i , with cdf F , and a threshold u . Compute excesses $X_i - u$. Interested in the distribution of the excesses when the threshold u increases.

Theorem: Balkema-deHaan(1974)-Pickands(1975). If the cdf $F^{[u]}(b_u + a_u x)$ of the excesses has a continuous limit as u increases, then the limit is one of three types of distributions: cdf's (in standard forms):



Exponential: $G(x) = 1 - e^{-x}$, $x \geq 0$,

Pareto: $G(x) = 1 - \frac{1}{x^\alpha}$, $x \geq 1$, $\alpha > 0$,

Beta: $G(x) = 1 - \frac{1}{(-x)^\alpha}$, $-1 \leq x \leq 0$, $\alpha > 0$.

For “light tailed” distribution of X_i , the limiting distribution of the excesses will be exponential.

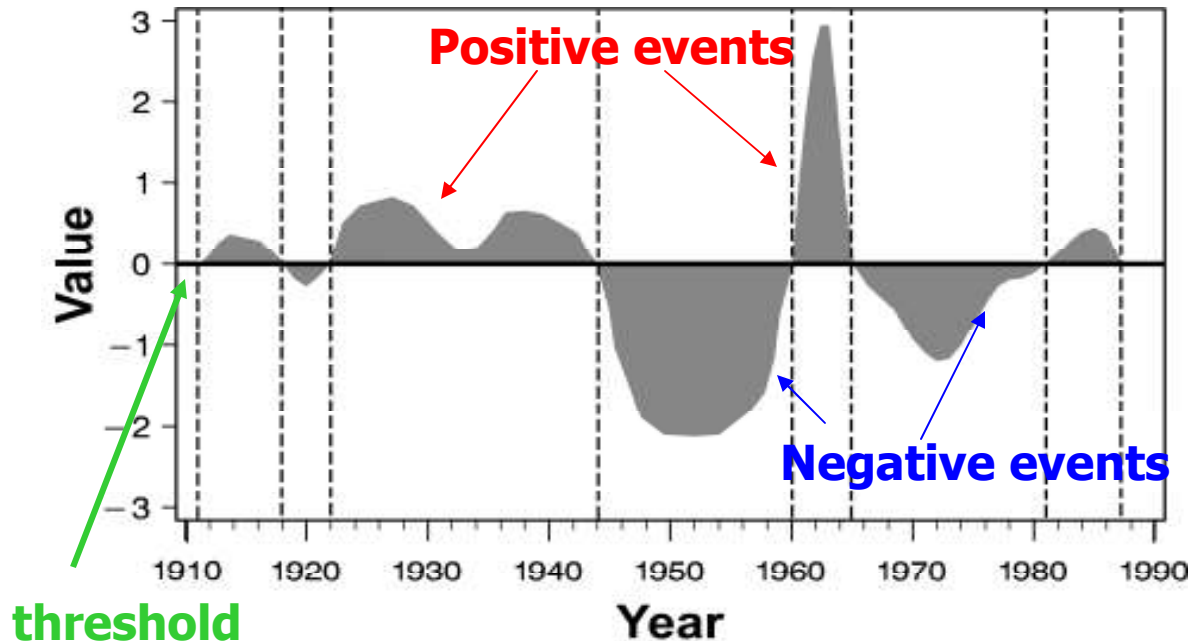
Summary up to now

Start with a process:

$$X_1, X_2, \dots, X_k,$$

DURATION = the number of years/time periods in one event : **random = N.**

$$\text{Duration} = N$$



MAGNITUDE = area of the shaded region = **random sum** of the series values for one event.

$$\text{Magnitude} = X = \sum_{i=1}^N X_i$$

MAXIMUM = max observation during an event = **random maximum** of the series values for one event.

$$\text{Max/Peak} = Y = \max_{i=1 \dots N} X_i$$

Random vector $(N, \sum_{i=1}^N X_i, \max_{1 \leq i \leq N} X_i) = (N, X, Y)$

GOAL: Construct a mathematically natural model for the JOINT distribution of (duration, magnitude X and maximum Y) of events.

Notable properties of the random vector (N, X, Y):

- All components are related/dependent, the joint behavior of X and Y is not trivial,
- The sum and maximum are of random number of random observations.

Hierarchical approach:

1. Specify distribution of N
2. Given N=n, **find** conditional distr. $f(x, y | n)$ of $(X, Y | N=n) = (\sum_{i=1}^n X_i, \max_{i=1, \dots, n} X_i)$
3. Get the joint distribution of (N, X, Y) as

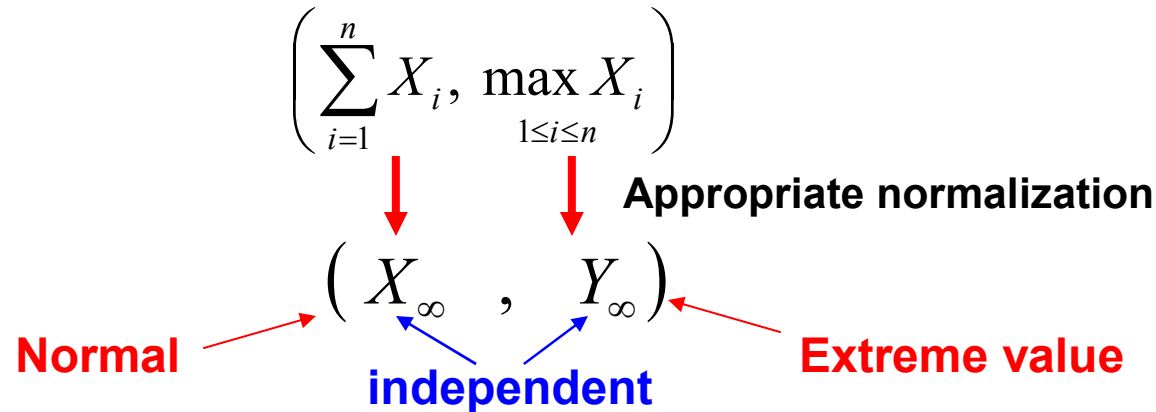
$$f(n, x, y) = f(x, y | n) f_N(n)$$

Joint Distribution of $(\sum_{i=1}^n X_i, \max_{i=1, \dots, n} X_i) = (X, Y)$

History: Start with X_1, \dots, X_n iid random variables, and known (nonrandom) n

GOAL: Find the joint distribution of (X, Y)

Only asymptotics for large n were available: X_i iid, X in sum DoA of Normal, Y in max DoA of some EV distribution, then



Main tool for limit theory is the hybrid characteristic -distribution function:

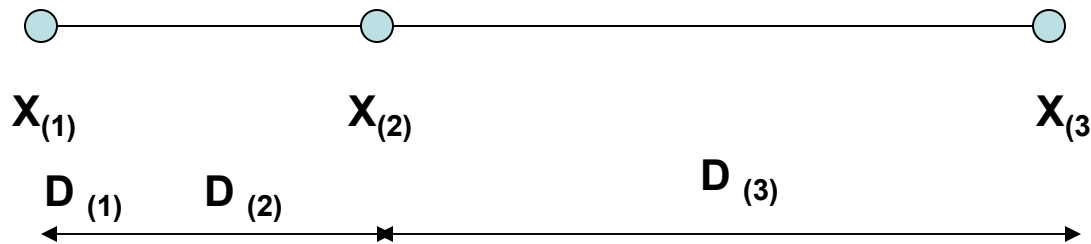
$$\psi_{(X,Y)}(t, v) = E\{(e^{itX} I(Y \leq v))\} = \left(\frac{1 - e^{(it-\beta)v}}{1 - it / \beta} \right)^n, \quad t \in R, \quad v \geq 0.$$

Intuitive understanding for the vector (X, Y)

X_1, X_2, X_3 iid $\exp(\beta)$

Consider $n=3$. Take order statistics $X_{(1)}, X_{(2)}, X_{(3)}$, and spacings

$D_{(1)} = X_{(1)}, D_{(2)} = X_{(2)} - X_{(1)}, D_{(3)} = X_{(3)} - X_{(2)}$.



$X = \sum X_i = 3D_{(1)} + 2D_{(2)} + D_{(3)}, Y = \max X_i = X_{(3)} = D_{(1)} + D_{(2)} + D_{(3)}$.

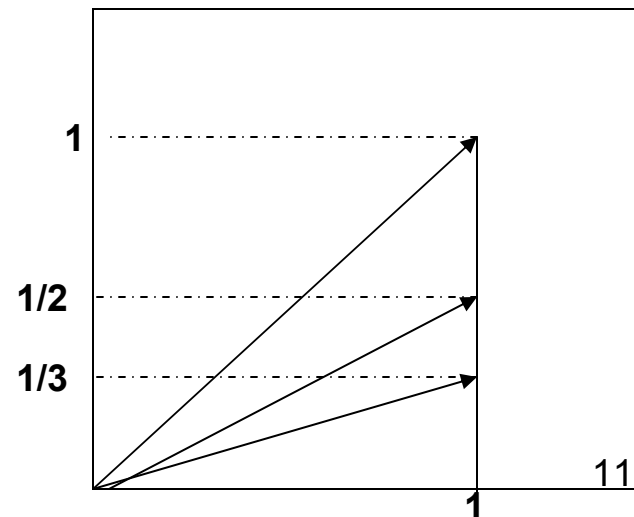
Take $W_i = (n-i+1) D_i$ iid $\exp(\beta)$, and

$D_i = W_i / (n-i+1)$ are independent exponential,

but not identically distributed.

$X = \sum X_i = \sum W_i, Y = \max X_i = \sum (W_i / i)$, so

$(X, Y) = \sum [W_i(1, 1/i)]$.



Stochastic Representation

Theorem. For any $n \geq 1$ and every $i = 1, 2, \dots, n$, let $a_i = (1, 1/i)$. Then a BGGE(β, n) random vector (X, Y) admits stochastic representation

$$(X, Y) \stackrel{d}{=} \left(\sum_{i=1}^n W_i, \sum_{i=1}^n (W_i / i) \right) = \sum_{i=1}^n a_i W_i,$$

No max,
only sums
in bivariate
case!

where the $\{W_i\}$ are IID exponential random variables with parameter β .

This shows that (X, Y) is infinitely divisible.

KEY NEW RESULT: PDF of the Bivariate Distribution of the Random Sum and Max for any n

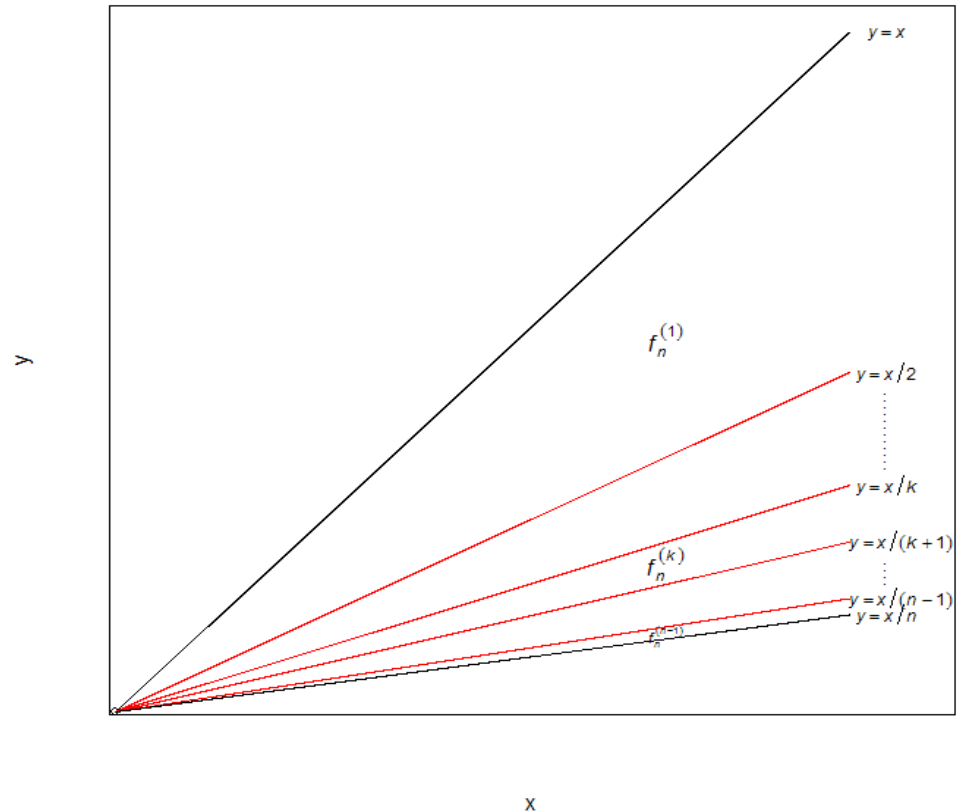
- **Domain** of the pdf

$$x/n \leq y \leq x$$

- PDF formula $f_n^{(k)}$ will depend on the **sector** k of the plane.
- There are **(n-1) sectors**:

$$S_0 = \{(x, y) : x=y>0\} \text{ and}$$

$$S_k = \{(x, y) \in R^2 : 0 \leq \frac{1}{k+1}x \leq y < \frac{1}{k}x\}, \quad k = 1, 2, \dots, n-1.$$



KEY NEW RESULT: PDF of the Bivariate Distribution of the Random Sum and Max for any n – contd.

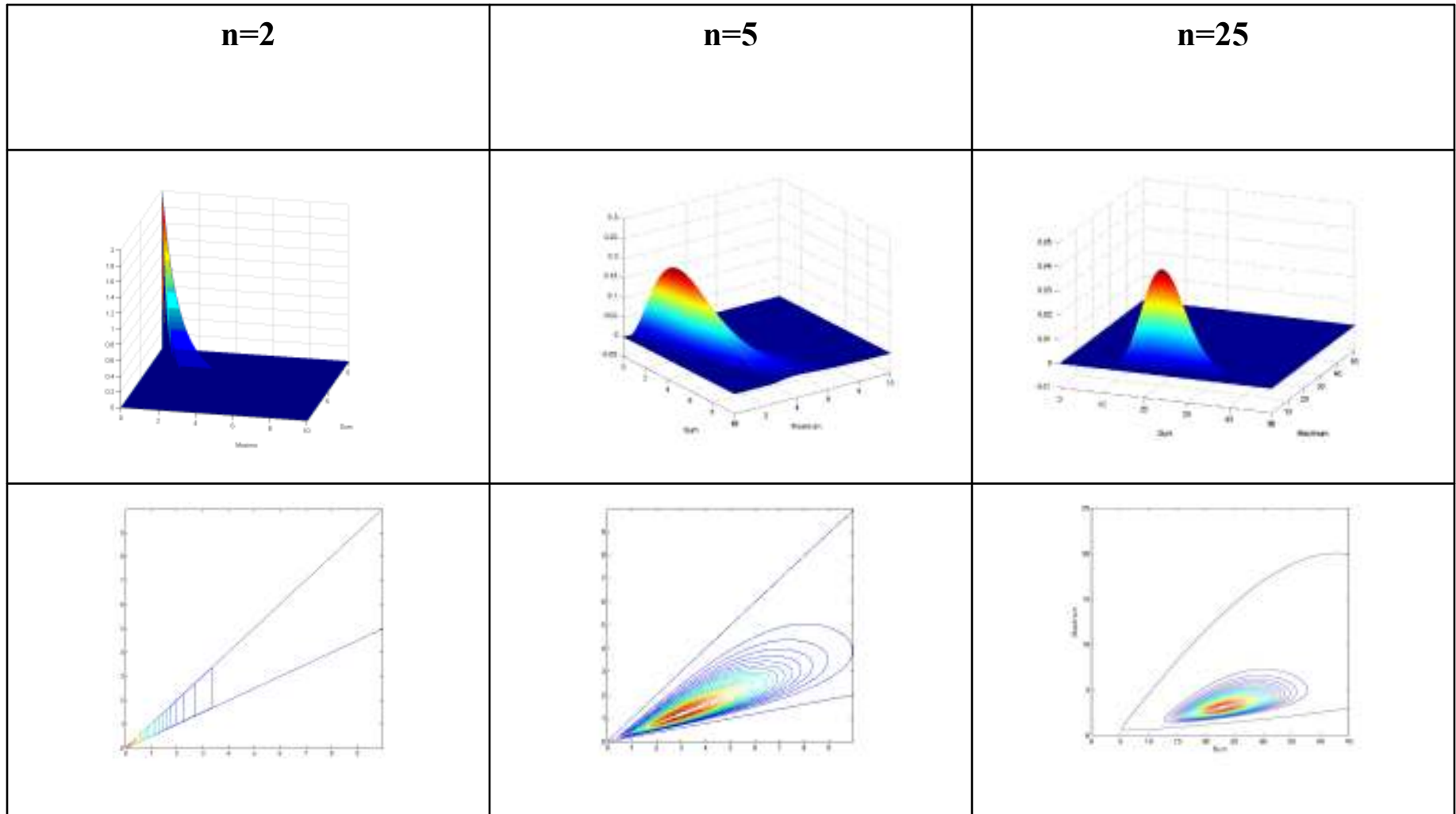
Let $X_i, i=1, \dots, n, n \geq 2$, be iid $\exp(\beta)$. The joint pdf of $(X, Y) = \left(\sum_{i=1}^n X_i, \max_{i=1, \dots, n} X_i \right)$ is given **explicitly** by

$$f(x, y) = \beta^n e^{-\beta x} H(x, y, n), \text{ where}$$

$$H(x, y, n) = \begin{cases} \sum_{s=1}^k \frac{n(n-1)}{(s-1)!(n-s)!} (x-sy)^{n-2} (-1)^{s+1} & \text{for } n \geq 2, (x, y) \in S_k \\ 1 & \text{for } n = 1, (x, y) \in S_0 \\ 0 & \text{otherwise.} \end{cases}$$

This new distribution is called **BGGE**(β, n) for **bivariate distribution with gamma and generalized exponential marginals**.

Densities



Marginal Distributions

Univariate

Marginal distribution of the sum X is $\text{Gamma}(n, \beta)$.

Marginal distribution of the maximum Y is generalized exponential $\text{GE}(n, \beta)$, (Gupta and Kundu, 2007) with cdf

$$F(y) = (1 - e^{-\beta y})^n, \quad y > 0.$$

Bivariate

Marginal distributions of (N, X) and (N, Y) .

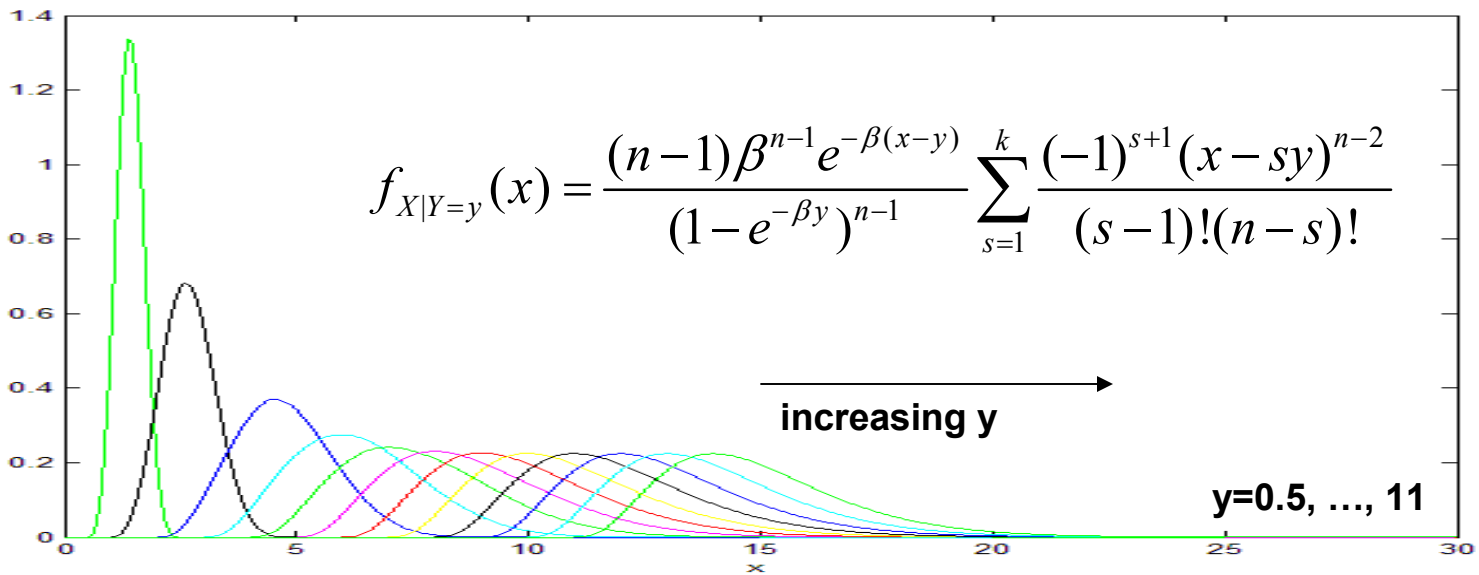
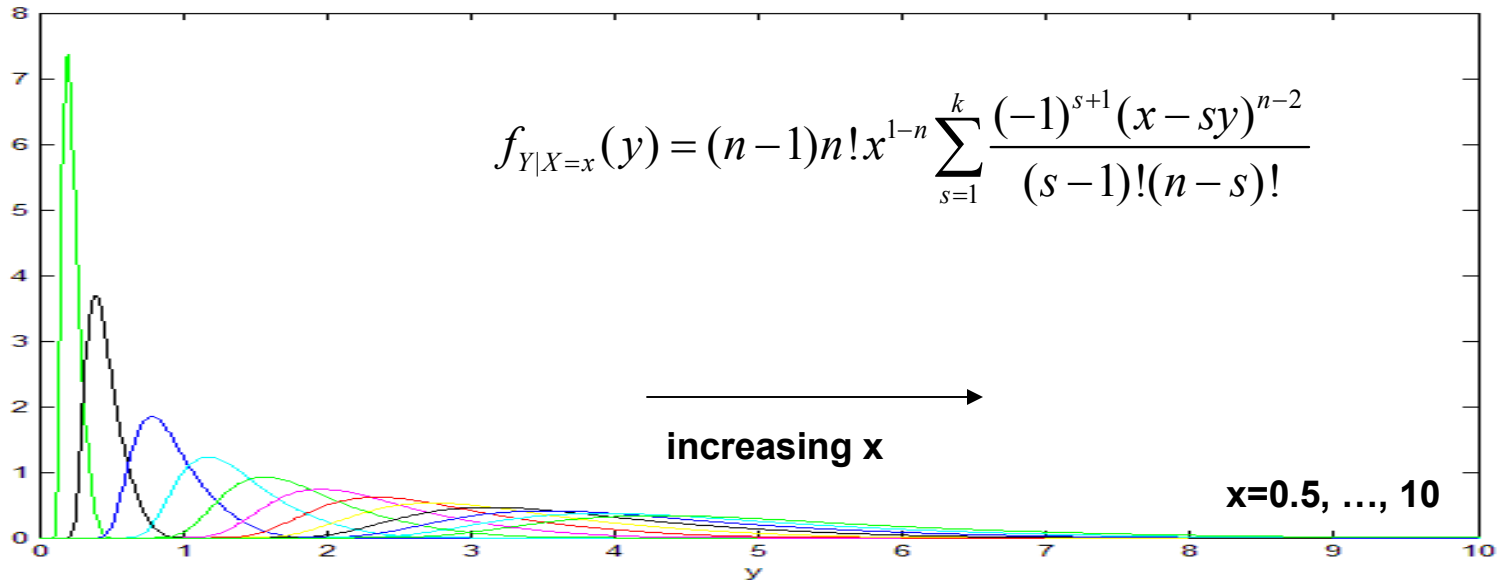
$$(N, X) = \underbrace{\left(N, \sum_{i=1}^N X_i \right)}_{\text{BEG}(\beta, p)}$$

BEG: Bivariate distribution with exponential and geometric marginals

$$(N, Y) = \underbrace{\left(N, \max_{1 \leq i \leq n} X_i \right)}_{\text{BTLG}(\beta, p)}$$

BTLG: Bivariate distribution with truncated logistic and geometric marginals

Conditional Distributions

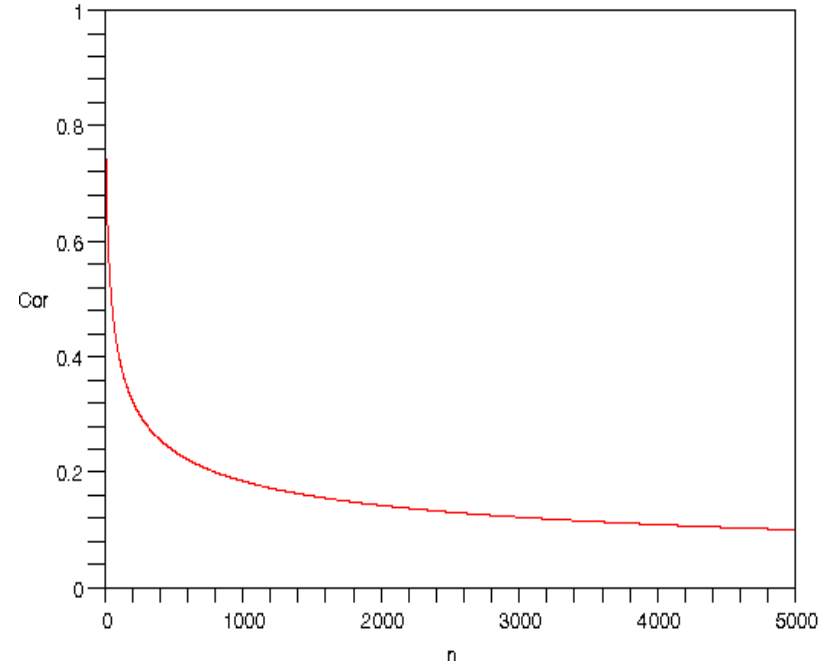


Covariance Structure

$(X, Y) \sim \text{BGGE}(\beta, n)$. $E X = n / \beta$, $E Y = (\psi(n+1) + \gamma) / \beta$, $\text{corr}(X, Y)$ is

$$\rho_n = \frac{\psi(n+1) + \gamma}{\sqrt{\frac{n\pi^2}{6} - n\psi'(n+1)}}$$

where ψ and ψ' are the digamma and polygamma functions, respectively, and γ is the Euler constant.



NOTE: $\rho_n \rightarrow 0$, but very slowly.

Thus, **approximation of (X, Y) with independent X and Y is not good even for large n .** **E.g. $\rho_{500} = 0.24$**

Distribution of the Ratio Y/X

History. Only limiting results were done for the ratio

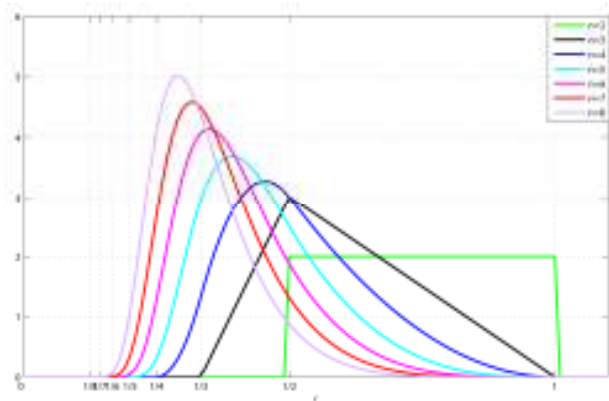
New Results.

Theorem. If (X, Y) is BGGE(β, n), then $(X, Y) \stackrel{d}{=} (X, RX)$,

where X and R are independent random variables, X has Gamma(n, β) distribution, and R has pdf of the ratio Y/X .

Theorem. Distribution of the ratio is parameter free, with density:

$$h_n(t) = n(n-1) \sum_{s=1}^k \binom{n-1}{s-1} (1-st)^{n-2} (-1)^{s+1}, \quad \text{for } \frac{1}{k+1} \leq t \leq \frac{1}{k}, \quad k=1, 2, \dots, n-1.$$



Morrison and Tobias 1965, O'Brien 1980, Maller and Resnick 1984, Haas 1992, Kesten and Maller 1994⁹

Trivariate Model: Joint Distribution of (X, Y, N)

N-random N-Geo(p), X_i 's iid $\exp(\beta)$ ind. of N.

$$(X, Y, N) = \underbrace{\left(\sum_{i=1}^N X_i, \max_{1 \leq i \leq N} Y_i, N \right)}_{\text{TETLG}(\beta, p)}$$

Trivariate distribution with **e**xponential, **t**runcated **l**ogistic and **g**eometric marginals¹.

The pdf of (X, Y, N) is $f(x, y, n) = \beta^n e^{-\beta x} p(1-p)^{n-1} H(x, y, n)$,

where function H is given in the pdf of the (X, Y|N=n)

¹ Kozubowski, Panorska, Qeadan (2011), JSPI, in press.

Estimation of Parameters

MLEs for β and p exist in explicit forms.

$(X_1, Y_1, N_1), \dots, (X_m, Y_m, N_m)$ random sample from TETLG(β, p) distribution.

The MLE's of β and p are:
$$\hat{\beta}_m = \overline{N}_m / \overline{X}_m \text{ and } \hat{p} = \frac{1}{\overline{N}_m}.$$

The vector MLE is consistent, asymptotically normal, and asymptotically efficient. The asymptotic covariance matrix is known.

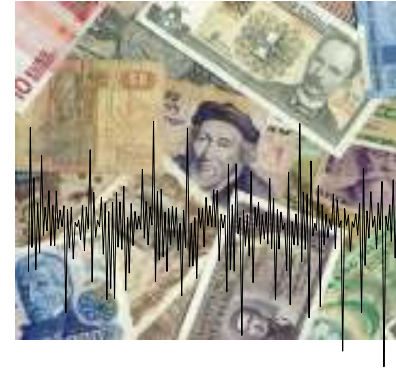
Note: The MLEs do not depend on Y_i 's.

Application 1: Financial Data

For the exchange rates data constructed data set of episodes.

Growth $X_i > 0$, **decline** $X_i < 0$

For **analysis**, considered episodes of **growth: $X_i > 0$** .



Our data were 1902 triples (X, Y, N) of growth episodes.

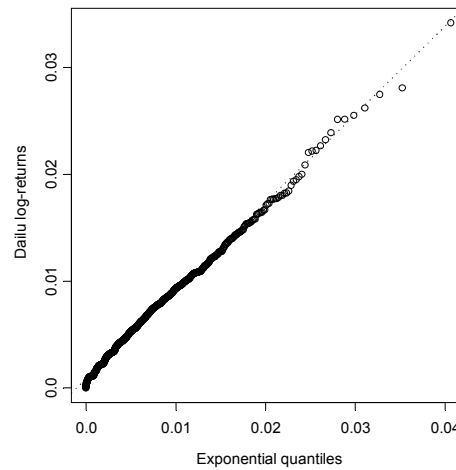
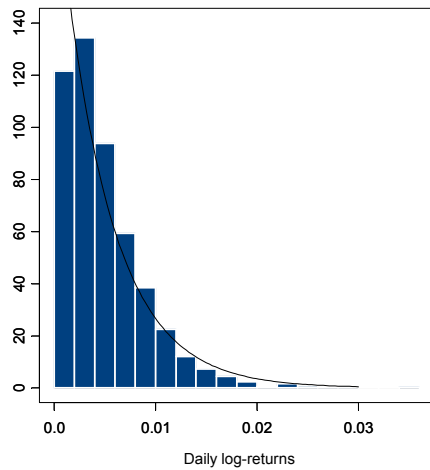
We checked that:

- Positive returns come from exponential distribution;
- Magnitudes of growth periods also come from exponential distribution;
- The fit of all bivariate marginals is quite reasonable;
- Fit of conditional distributions is reasonable

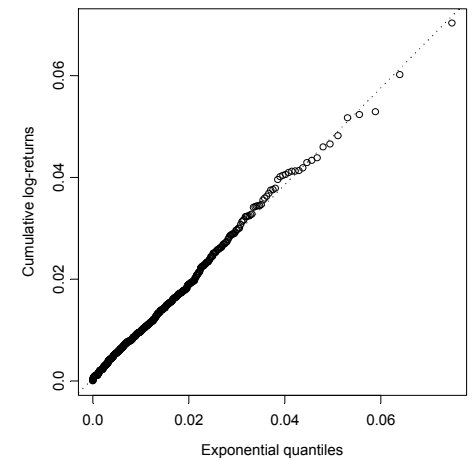
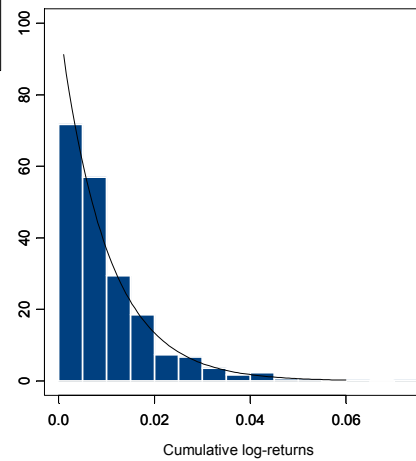
Stability of the returns wrt growth periods: Cumulative positive returns have the same distribution as the returns.

Modeling financial data

Daily positive log-returns X_i fit with exponential distribution

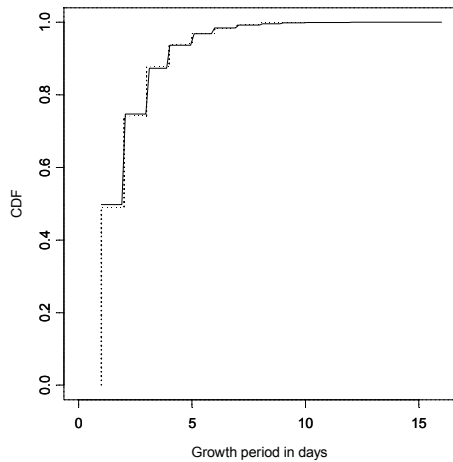


Cumulative positive log-returns $X = \sum X_i$ fit with exponential model

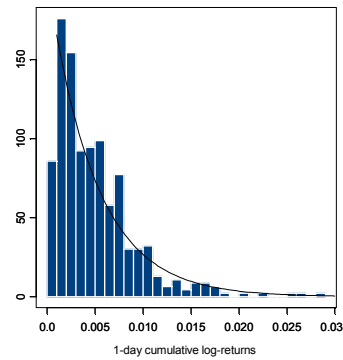


Modeling financial data: Fit of BEG model to $(N, X = \sum X_i)$

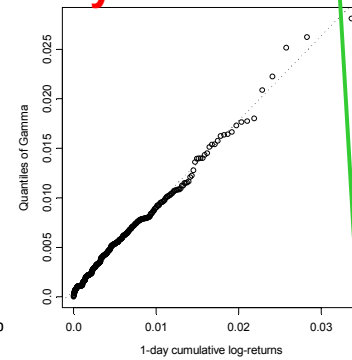
Geometric fit to durations



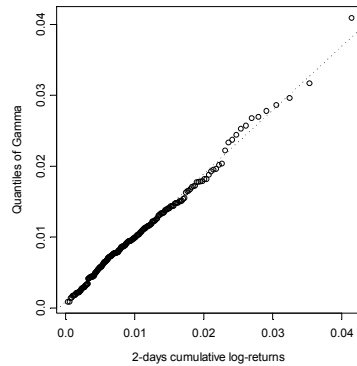
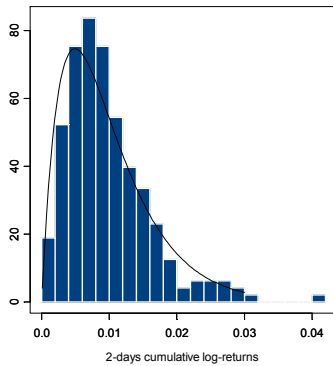
Gamma fit to the conditional distributions



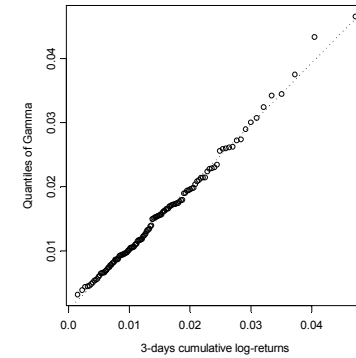
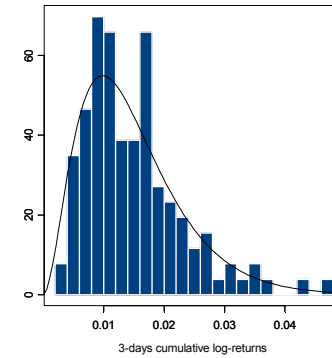
1-day events



2-day events



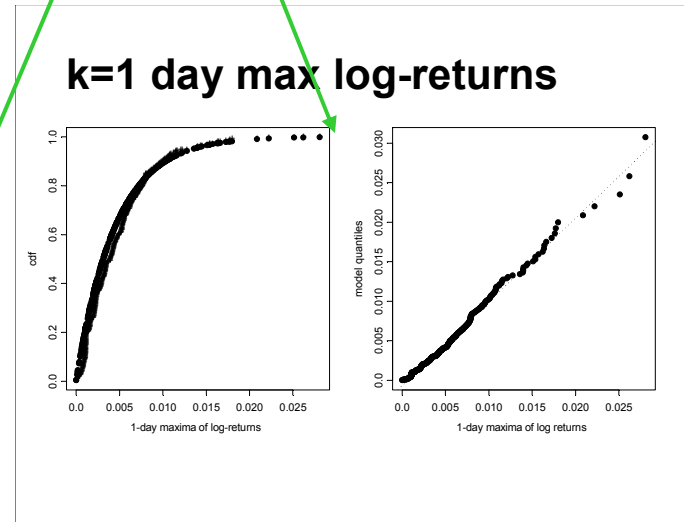
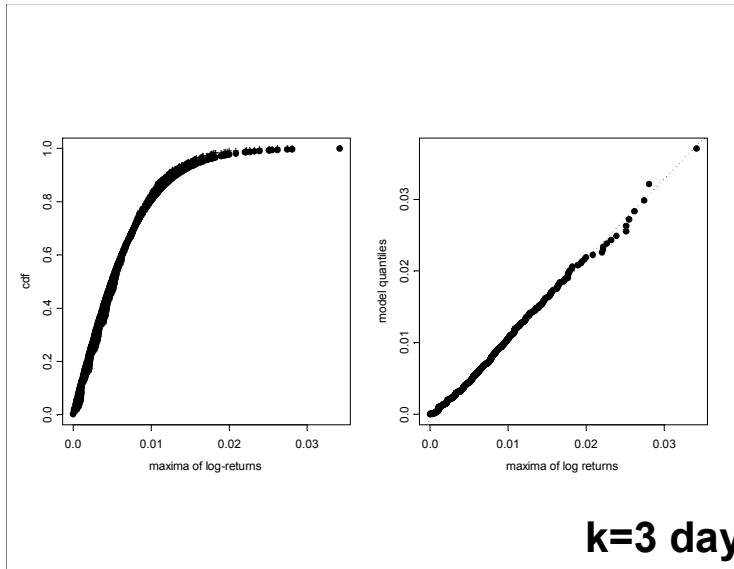
3-day events



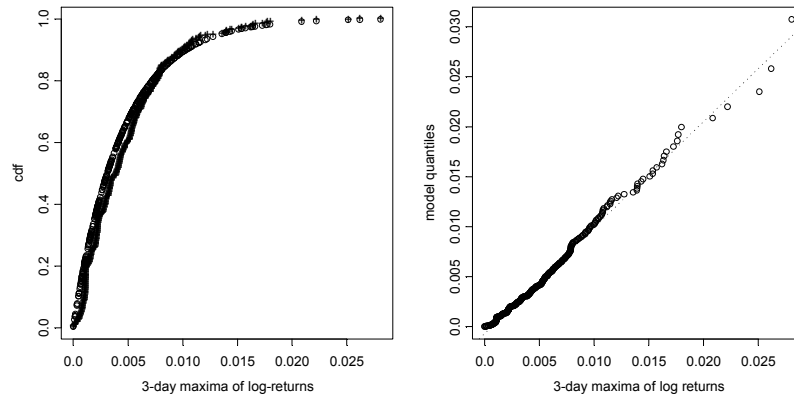
Modeling financial data: Fit of BTLG model to $(N, Y=\max X_i)$

Marginal distribution of the maximum fit.

Fit of conditional distr. $X|N=k$



k=3 day max log-returns



Fit of TETLG model to Exchange data

All tests on 0.05 significance level. Checking bivariate fit.

Is the ratio independent from the sum?

Checked Pearson's correlation between ratio and sum;

Ho: $\rho=0$ Ha: $\rho\neq 0$

$r = 0.179$, p-value = 0.1749

Conclusion: Ratio and sum are not correlated, so BGGE model OK

Bivariate Kolmogorov-Smirnov goodness of fit test.

Ho: data comes from BGGE distribution **Ha:** data does not come from BGGE

Test stat=1.346, p-value=0.45

Conclusion: BGGE model OK

Bivariate Kolmogorov-Smirnov goodness of fit tests for conditional distributions of (X, Y) GIVEN $n=N$.

Test results below.

N=2	3	4	5
P-value=0.293	0.806	0.335	0.770

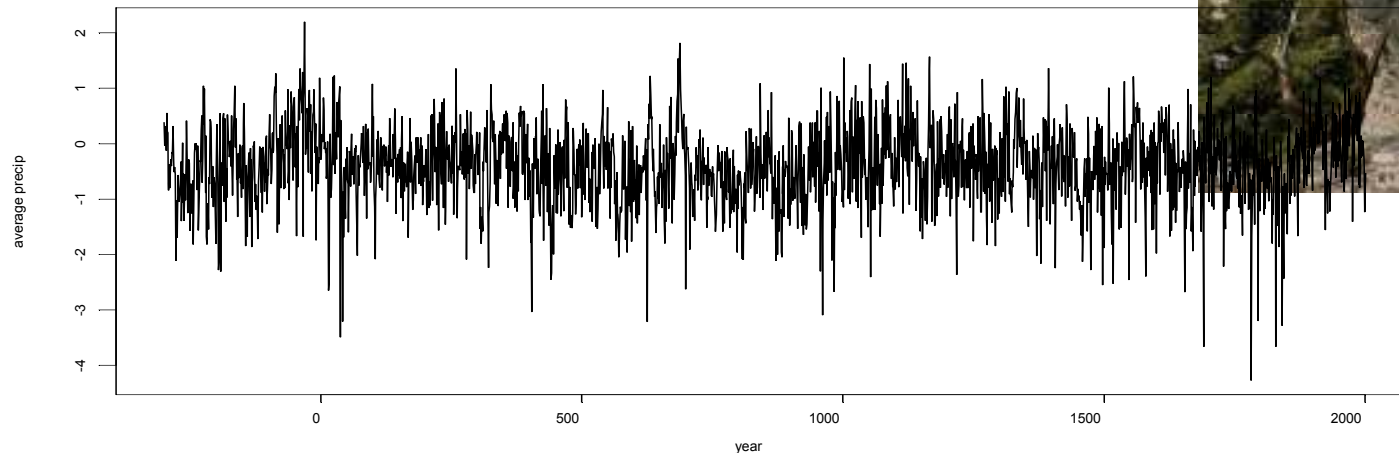
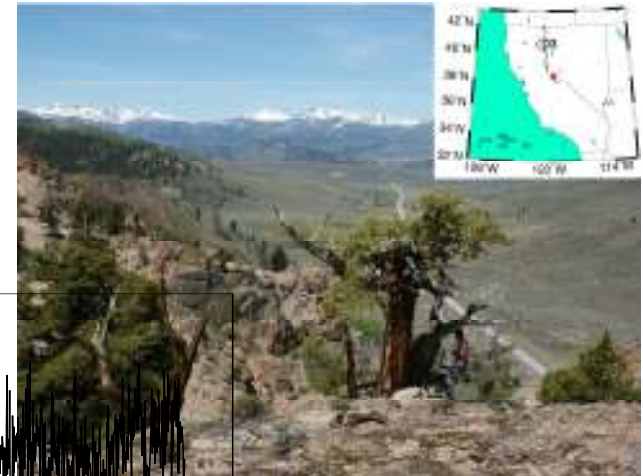
OVERALL CONCLUSION:
REASONABLE FIT of all the bivariate models (BEG, BTLG, and BGGE) for the growth episodes

Example 2: Climate and Hydrology: The “Dust Bowl”

The great drought of the 1930s in the US (and Canada), the 'Dust Bowl' impacted millions of people in many states. It had terrifying effects on the economy and the natural environment.

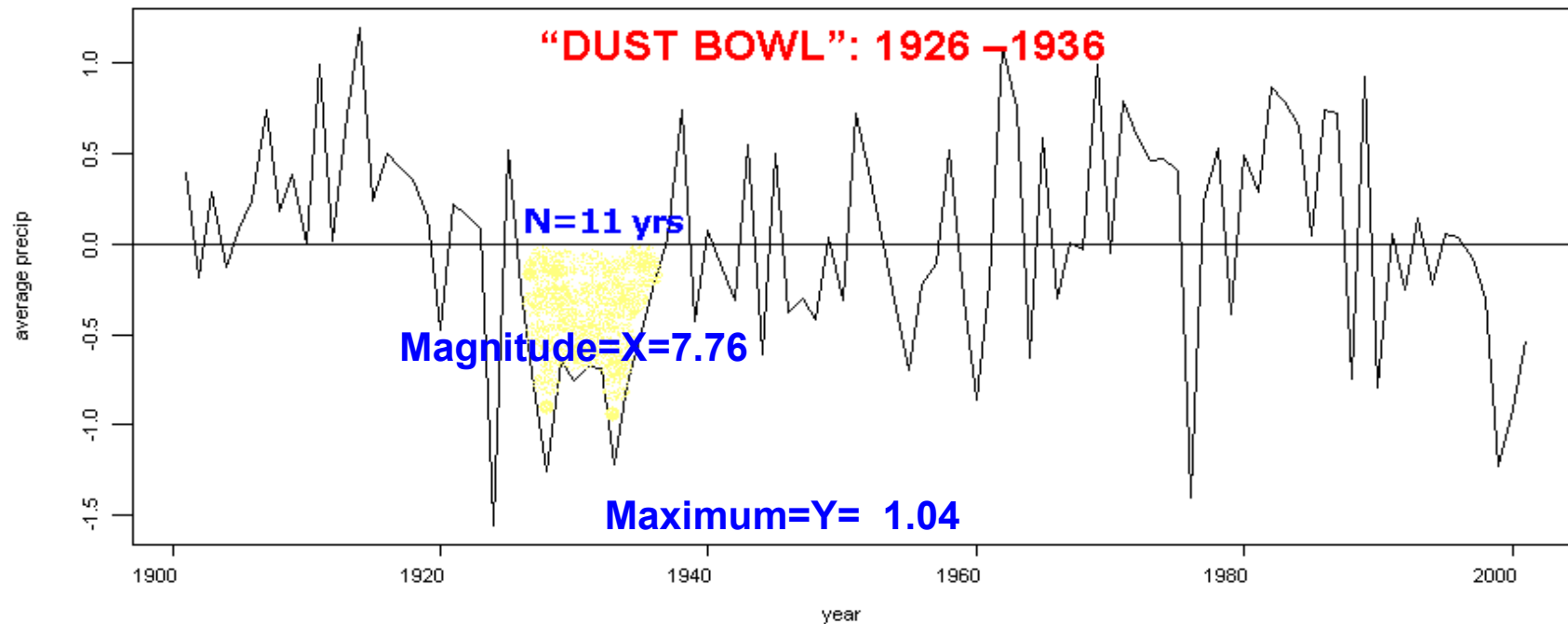


The data: Dendroclimatic (western juniper) reconstruction of precipitation from 300 BC to AD 2001 in the Walker River watershed (California/Nevada). California Climate Division 3.



Total annual precipitation (full reconstruction) plotted as deviation from the overall median in standard deviation units.

EXAMPLE 2: THE “DUST BOWL” contd.

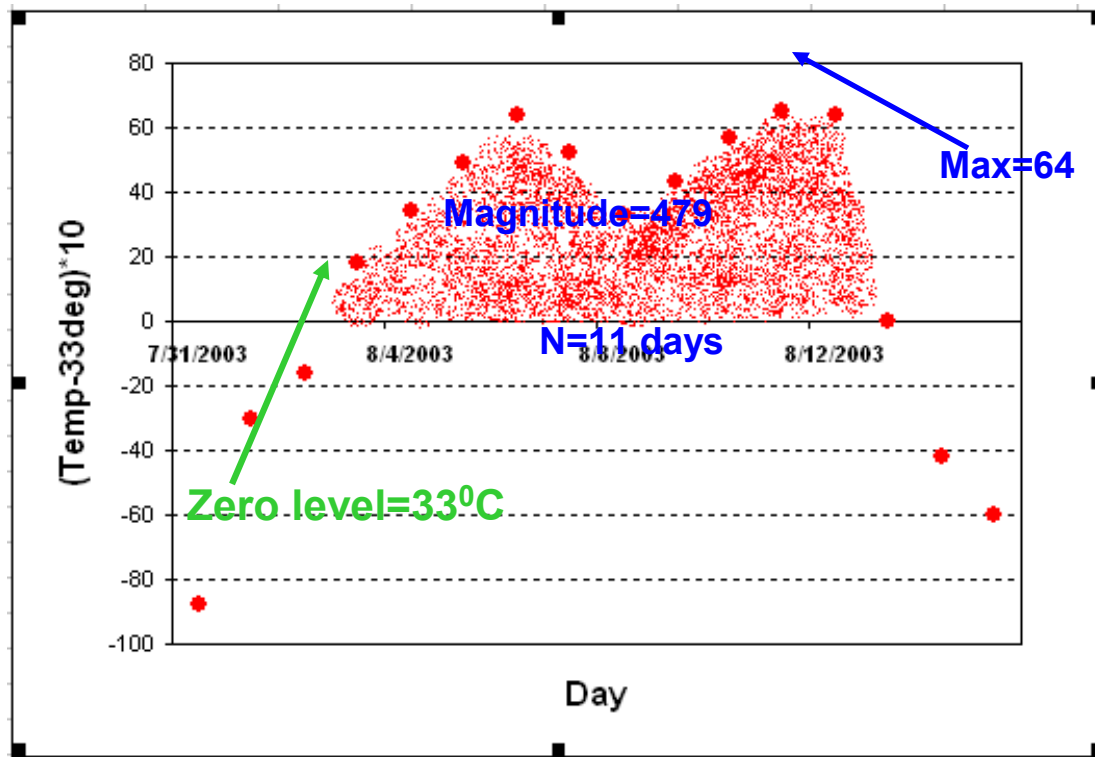


- Probability of a drought longer or larger than the 'Dust Bowl' is 0.08;
- Probability of a drought longer and larger than 'Dust Bowl' is 0.06.
- Conditional probability of a drought with at least 'Dust Bowl's' magnitude given that duration is 11 years is 0.46

EXAMPLE 3: Paris heat wave of the 2003

In August 2003, France experienced an extreme heat wave, that resulted in an estimated 14,802 deaths*.

Definition of a hot event: consecutive observations above the 33°C.



•Probability of a hot event with the same duration of 11 days is 0.000075;

•Conditional probability of a heat wave with at least that magnitude given that duration is 11 days is about $5.5e-4$.

•Probability of a heat wave longer than 6 days and larger than 100 (98th percentile of magnitudes) is 0.005.

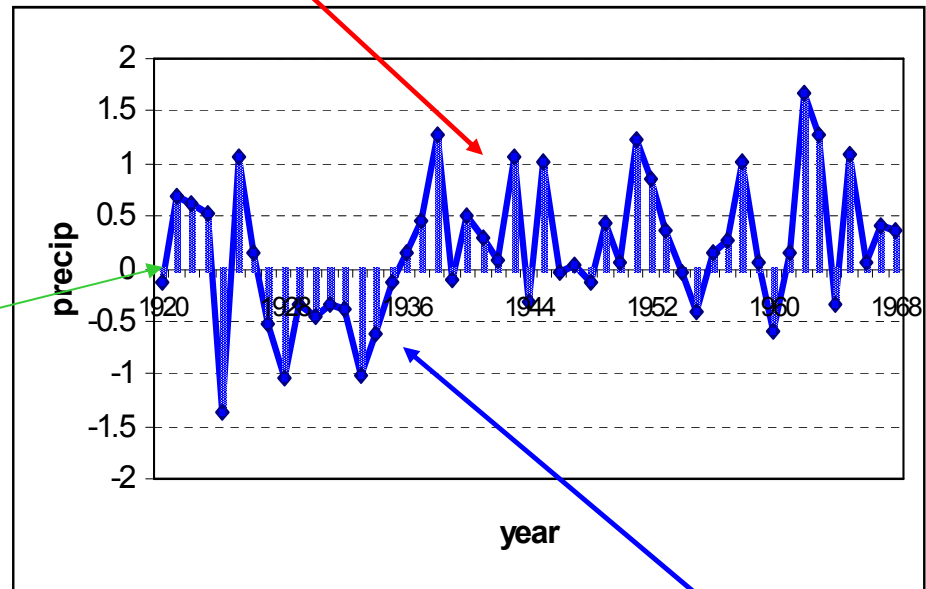
Note: Maximum Y = 64deg C*10 (really 39.4oC)

SUMMARY

Given a **process** X_1, X_2, X_3, \dots **fluctuating around a threshold.**

Suppose X_i 's are iid $\exp(\beta)$,
and the process
fluctuates according to a
geometric model.

Positive events



threshold

$$(X, Y, N) = \underbrace{\left(\sum_{i=1}^N X_i, \max_{1 \leq i \leq N} Y_i, N \right)}_{\text{TETLG}(\beta, p)}$$

Negative events

NEW stochastic model for the joint distribution of duration, magnitude,
and maximum of the events.

WHY IMPORTANT FOR NEVADA and OTHER WESTERN STATES ?

We have to deal with lots of extremes.

NEVADA: A STATE OF EXTREMES



Thank You